

AN ESTIMATION OF CHINESE RENMINBI EXCHANGE RATE IMPACT ON THE REAL EXPORTS OF INDONESIA TO THE US: IS THERE A J-CURVE?¹

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ABSTRACT

This paper seeks whether a J-Curve exists on the impact of changes in the Chinese Renminbi (RMB) exchange rates on bilateral exports of Indonesia to the United States (US), particularly in the long run. The Johansen cointegration procedures and Vector Error Correction Model (VECM) regression are applied. The cointegration test shows that there are long-term relationships amongst real GDP of US, Indonesian Rupiah (IDR), real exchange rates and volatility, and Chinese RMB real exchange rates. The result shows that the RMB exchange rate has a negative significant impact (substitution relationship) on Indonesian export to the US. The result also suggests a dissatisfaction of the Marshall-Lerner condition indicating the J-curve phenomenon does not exist.

Keywords: RMB, Marshall-Lerner Condition, J-Curve, Johansen Cointegration, VECM

JEL Classifications: G15

INTRODUCTION

In recent years, the world is woken up by the fact that China has become an economic giant. Ishihara (1993) and Panagariya (2006) recorded that the Chinese economy grew up to around 8-10% after the market had been expanded. Cable and Ferdinand (1994), World Bank (2006) and Bloom et al (2006) argued that one of the main reasons for Chinese economic succession is the support of their government in issuing policies related to the openness of international trade. In this context, Baak (2005, 2007) and Goldstein and Lardy (2006) stated that the key factor of the China's succession today is mainly caused by the pegged of Renminbi (RMB). It was pegged at 8.704 against US dollar (US\$) after the first devaluation in 1994. Many economists then inferred also that China has already started the currency war.

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Funke and Rahn (2005) warned that the main purpose of RMB devaluation is to maintain Chinese export relatively cheaper compared to others countries in international market. This policy is basically aimed to increase Chinese current account. Appleyard, Field and Cobb (2011) and Krugman and Obstfeld (2006) clarified further that eventhough a real currency devaluation exacerbates domestic current account at the beginning, this policy may potentially improve their current account after some months later, assuming a Marshall-Lerner doncition holds.

The main problem of RMB devaluation is happened in its side effect. There are many other countries in practice suffering by RMB devaluation policy. Funke and Rahn (2005) noticed that after RMB had been devalued it had lead to several job losses in the US, Japan and other economies in the East Asian region. Baak (2007, 2005) observed that the depreciation of RMB against the US dollar had turned out to decrease both in Korean exports to Japan and the U.S exports to China. Many researchers then warned that China has also become the major competitor country in the world market (Teh, 2009).

In practice, as mentioned by Widodo and Putriani (2011) and Putriani (2010), China has played roles as one of major trading partner for Indonesia. Considering these empirical evidences, this paper aims to seek the existence a J-curve and Marshall-Lerner condition on the RMB real exchange rate and to estimate its impact in terms of bilateral exports relationship between Indonesia to the United States (US), particularly in the long-run.

This study will only be focused on the impact of Chinese Renminbi on total bilateral real exports between Indonesia to the US in the long run from demanded exports sides. Therefore, the analysis for estimated coefficient of real GDP, real exchange rate and volatility (RERVOL), supplied import function and sectoral (commodity) level analysis will be excluded.

The rest of this paper is organized into five parts. Part 2 exhibits the literature review. Part 3 examines research method along with the econometric model and data. Part 4 shows the empirical results and analysis. Some conclusions are presented in Part 5.

LITERATURE REVIEWS: AN OVERVIEW OF MARSHALL-LERNER CONDITION AND J-CURVE PHENOMENON

Jamilov (2010), Husman (2005), and Kapoor and Ramakrishnan (1999) explained that a change in exchange rate has two effects on trade flow, namely price effect and volume effect. The price effect means that in the short run, the depreciation will cause the price of imports more expensive and the price of exports cheaper for foreign buyers. Figure 1 shows that if the volume of imported and exported goods are not significantly change, the trade balance will gradually go down (the movement from point 1 to point 2). Since the goods and services of exports and imports are sufficiently elastic, the volume of trade will quickly response to the depreciation and the current account improves (the movement from point 2 to point 3). Krugman and Obstfeld (2006) summarized that if there is a domination of the volume effect over the price effect, the phenomenon of Marshall-Lerner condition will be satisfied. The total effect when the plotted over time with trade balance on the y-axis will yield the J-curve.

Considering the objective of this research, the Marshall-Lerner condition can be mathematically detected by summing up the elasticity of export demand (η) and elasticity of import demand (η^*). Following Krugman and Obstfeld (2006), the condition for an increase in price to improve the current account, i.e. Marshall-Lerner condition, can be formulated as follows:

$$\eta + \eta^* > 1 \quad (1)$$

Equation (1) implies that if current account is initially zero, a real currency depreciation (devaluation) leads a significant improve in current account. The surplus of current account will be reached since the sum of the relative price elasticity of export and price elasticity of import demand exceeds 1 (one). Considering equation (1), this study will only employ one model of export equation. Since the value of estimated equation is positively greater than 1, it implies that the Marshall-Lerner condition and J-Curve phenomenon might be observed.

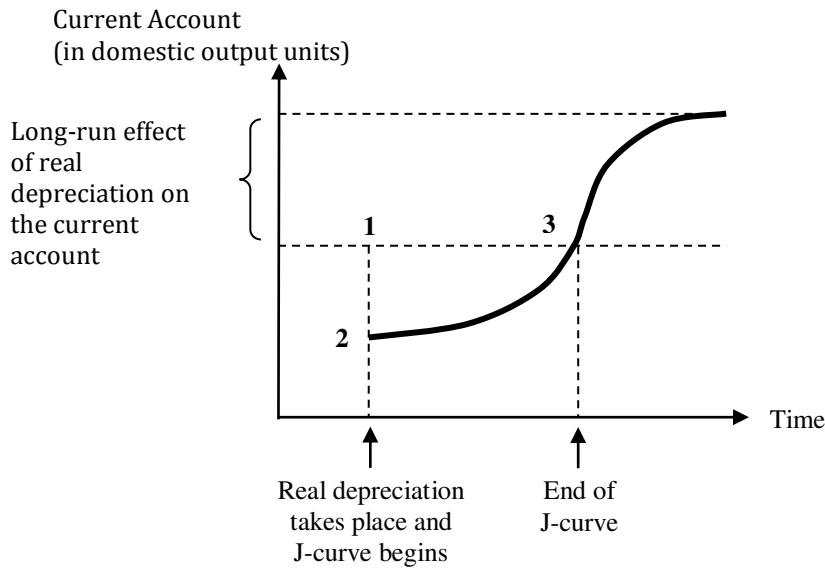


Figure 1. The J-Curve
Source: Krugman and Obstfeld (2006)

RESEARCH METHODS

Cointegration Test

As pointed out by Gujarati (2003), the existence of cointegrating relationships is indicated when two or more non stationary series have at least a linear combination which is stationary, $I(0)$. In short, the stationary linear combination cancels out the stochastic trends in two series, proven by checking the residuals from the regressions which are stationer, $I(0)$. Thomas (1997) noticed that the presence of cointegration for multivariate equation case is detected by applying Johansen cointegration test. The empirical result of Johansen test will be confirmed by Error Correction Model (ECM), as follows:

$$\begin{aligned} \Delta EX_{ijt} = & \alpha + ECT_{ijt-1} + \sum_{h=0}^{nx} \beta_h \Delta EX_{ijt-h-1} + \sum_{h=0}^{np} \gamma_h \Delta RER_{ijt-h} + \\ & \sum_{h=0}^{np} \delta_h \Delta GDP_{jt-h} + \sum_{h=0}^{np} \eta_h \Delta RERVOL_{ijt-h} + \\ & \sum_{h=0}^{np} \varphi_h \Delta RERC_{cjt-h} + u_{ijt} \end{aligned} \quad (4)$$

Here, the lengths of involved for each variable is denoted as nx , np , ns , and nc . As pointed out by Baak (2005, 2006, and 2007), the negative and significant of the estimated error correction term (ECT) coefficient (ECT_{ijt-1}) represents the presence of long-run relationships among the variables included in the model. Since there is at least one cointegration among the variables, the causal relationship among these variables will be detected by the VECM procedures.

Data Sources

This study takes sample period from the first quarter of 1987 to the first quarter of 2011. All of calculated data were taken from *Direction of Trade Statistics (DOTS)* of the International Monetary Fund (IMF), *International Financial Statistics (IFS)* of the IMF and *National Bureau of Statistics of China (NBSC)*.

Hypotheses Development

The dynamics of aggregate demand for exports is constructed within the basic theory of demanded exports. Specifically, as mentioned by Murianda (2008), the demanded exports (X_d) depends on price of exports goods and services (P_x), real income of importing country (Y_f) and price of exports goods and services from competitor (P_c). The demanded exports function can be constructed, as follows:

$$X_d = f(P_x, Y_f, P_c) \quad (3)$$

Equation (4) shows the econometric model of how real exchange rates of RMB affects bilateral exports of Indonesia in the long run. It is constructed by referring to the equation 1 and some previous model such as Baak (2005, 2006 and 2007), Arize and Osang (2000), Widarjono (2005), and Funke and Rahn (2005). In addition, it is augmented dummy variables in order to catch up the impact of RMB devaluation in 1994.

$$EX_{ijt} = \beta_0 + \beta_1 GDP_{jt} + \beta_2 RER_{ijt} + \beta_3 RERVOL_{ijt} + \beta_4 RERC_{jt} + \beta_5 D_{ijt} + \varepsilon_{ijt} \quad (4)$$

All of variables are measured in natural logarithm. Subscript i represents exporter (Indonesia) and j denotes importer (US), while c symbolizes China takes part as a competitor of Indonesia in the market of country j . Subscript t represents time of quarterly data. EX_{ijt} , the real exports of i to j is measured for representing the true value of

demanded exports. Following the work of Baak (2006), the formula for, EX_{ijt} , can be formed by simple formula as appeared in appendix.

As a proxy to measure the real income of one country, variable GDP is involved in the model, GDP_{jt} . Data for real GPD are provided by the International Financial Statistics (IFS). It is not clear that GDP and real export has positive or negative relationship. If the estimated coefficient sign is negative, it means that demanded goods are inferior. Meanwhile, if the estimated coefficient sign is positive, it suggests that goods demanded are normal. However, when the estimated coefficient elasticity is greater than 1, it indicates that goods demanded are superior.

Real exchange rate, RER_{ijt} , represents i country currency against j 's currency. It represents the price of demanded exports. It is computed as common formula of RER in many literatures of international trade. The effect of expected coefficient of bilateral RER on the bilateral exports estimation can be significantly positive or negative. When exchange rate of exporting country depreciates and the estimated coefficient is positive, it reflects the elastic export commodity. Meanwhile, when the exchange rate of exporting country depreciates and the estimated coefficient is negative, it represents the export commodity which is inelastic. The Marshall-Lerner condition will be satisfied since the estimated coefficient of RER is greater than 1, implying that depreciation will improve the current account and detect the existence J-curve, and vice versa. In addition, in the case of Chinese real exchange rate, $RERC_{cjt}$, the measurements is also formulated in the same way by converting the subscript i into c and the interpretation of its estimated coefficient is similar with RER_{ijt} variable.

Meanwhile, the $REXVOL_{ijt}$ represents the logarithm value of real exchange rate volatility. It is measured as natural logarithm of the absolute quarterly standard deviation of monthly real exchange rate (see appendix). D_{ijt} denotes dummy variable used to investigate the impact of RMB devaluation on the real exports of Indonesia to US, whereas for periode before devaluation (1987.Q1-1993.Q4) it gives a value of 0 (zero), while for the period after devaluation (1994.Q1-2011.Q1) it gives a value of 1(one).

EMPIRICAL RESULTS AND ANALYSIS

Unit Root Test

The equilibrium relationship is detected by testing the order of integration of the variables examined i.e. real exports, real GDP, RER, RER volatility, and RER China. To do so, we firstly apply the Philip-Perron (PP) test to see for both the level and the first difference. The optimal length of lag applied in the test is determined by using Aikake Information Criterion (AIC). As presented in table 1, all variables of PP statistics are lower than MacKinnon critical value, meaning that all variables are not stationary at level. However, all variables are integrated of order one I(1) indicating that they are stationary at the first difference.

Table 1: PP Unit Roots Test for the Level and the First Differences: Indonesia-US

Indonesia to the US	Level			First Difference		
	Lag	PP Test	AIC	Lag	PP Test	AIC
Export						
Intercept	1	-0.656	-5.412	2	-9.179 *(I)	-5.399
Trend and Intercept	1	-1.829	-5.420	3	-9.158 *(I)	-5.379
None	3	4.249	-5.424	3	-7.855* (I)	-5.267
Real GDP						
Intercept	1	-2.201	-7.228	3	-5.650 *(I)	-7.43
Trend and Intercept	3	-0.320	-7.218	3	-5.938 *(I)	-7.435
None	3	6.078 *	-7.157	1	-3.546 *(I)	-7.308
RER						
Intercept	2	-2.220	-1.724	2	-7.840*(I)	-1.723
Trend and Intercept	2	-2.151	-1.703	2	-7.836*(I)	-1.706
None	2	-0.046	-1.708	2	-7.880*(I)	-1.744
RER Volatility						
Intercept	1	-2.145	3.457	3	-17.035*(I)	3.324
Trend and Intercept	3	-5.578*	3.457	3	-16.932*(I)	3.345
None	1	-0.434	3.498	3	-36.119*(I)	3.306
RER China						
Intercept	1	-2.407	-3.180	2	-7.584*(I)	-3.162
Trend and Intercept	3	-0.734	-3.166	1	-8.044*(I)	-3.192
None	1	-1.547	-3.101	3	-7.237*(I)	-3.143

Source: IFS, DOTS and NBSC. Author's calculation.

The figures in the brackets are the order of integration

*) denotes rejection of a unit root hypothesis based on Mackinnon's critical values at the level of significance $\alpha=1\%$, 5% , or 10% .

Cointegration Test

The optimal lags length in Johansen cointegration test are chosen based on the VAR lag order selection criteria tests, the which maximum lags length included in the tests are 1 (one) (see Appendix 1). In addition, if the I(1) variables involved are cointegrated, thus the estimated coefficient are consistent (Thomas, 1997).

Table 2. Johansen Cointegration Test

Null Hypothesis	Test Statistic		Critical Values (5%)	
	Trace	Max	Trace	Max
$r = 0$	93.714	46.984*	107.346	43.420
$r \leq 1$	46.730	21.361	79.341	37.164
$r \leq 2$	25.368	10.954	55.246	30.815
$r \leq 3$	14.414	7.568	35.010	24.252
$r \leq 4$	6.846	6.387	18.398	17.148

Source: IFS, DOTS and NBSC. Author's calculation.

*) denotes rejection of a unit root hypothesis based on Mackinnon's critical value at at the $\alpha=1\%$, 5%, or 10%.

As presented in table 2, the result shows that both trace statistics and max-eigenvalue statistics confirm the presence of cointegrating vectors, implying the variables in equation (4) are cointegrated. This finding is supported by the Error Correction Model (ECM) estimation presented in table 3 showing that the estimated coefficient of Error Correction Term (ECT) in all of regression results are negative and significant at the level of significances of $\alpha= 5\%$ and 10%. In other words, this result shows long-run relationships among variables involved.

Table 3: ECM Estimation

C	Δ GDP	Δ RER	Δ RERVOL	Δ RERC	DUM	ECT	R ²	DW
0.001 (0.002)	0.96* (0.220)	-0.07 (0.014)	0.002* (0.001)	-0.06 (0.042)	-0.07 (0.021)	-0.12* (0.045)	0.27	1.87

Source: Authors' calculation.

Notes: Standard error in parantheses

The asterisk (*) indicate the rejection of the null hypothesis of a zero coefficient at the 5% significance level.

The Estimated Coefficient Interpretation

As mentioned before, this study only focuses on the impact of Chinese Renminbi on bilateral exports of Indonesia to the US in the long run, hence the short run impacts will be ignored. Equation (5) shows, the Indonesian exports are only positively influenced by GDP and bilateral real exchange rate of IDR against the US\$ in the long run. The coefficient value of GDP of the US informs that offered products by Indonesian are superior product for the US people. Subsequently, 1% depreciation of IDR will turn out 0.32 percent increase in exports goods from Indonesia to the US. Since the estimated coefficient of real exchange rate is not greater than 1, the Marshall Lerner condition does not exists in the model, implying that the depreciation of IDR might not improve the trade balance.

$$EX_{ijt} = 4.27 - 1.52GDP_{jt} + 0.32 RER_{ijt} - 0.10RERVOL_{ijt} - 0.36RERC_{cjt} - 0.10 D_{ijt} \quad (5)$$

[3.45]* [3.77]* [7.32]* [4.37]* [1.57]

The asterisk (*) indicate the rejection of the null hypothesis of a zero coefficient at the 5% significance level.

Both the bilateral real volatility exchange rate of IDR against the US\$, RERVOL, and bilateral real exchange rate of RMB against US\$, RERC, in contrast, has negatively significant impact to the bilateral export of Indonesia to the US. It means that the exporters of Indonesia are risk averse and the relationship of exports goods between Indonesia and China in the US market is substitution, respectively. These findings are similar to Baak (2007) that also found the relationship of export goods between Korean and China in the Japan market. In contrast, Widodo and Putriani (2011) and Putriani (2010) observed that the relationship between Indonesia and China applying OLS-ECM procedures, showing complementary. The devaluation policy of RMB denoted by the dummy variables, D_{ijt} . In practice, it have no significant impact on the bilateral exports of Indonesia to the US.

Impulse Response Function and Variance Decomposition

Enders (1995) suggests the impulse response analysis and variance decompositions to define the relationships amongst variables dependen and independen. Husman (2005) further noted that the adjustment of trade balance dynamic against shock can be observed

by measuring its impulse response function. Figure 2 shows the impulse response function concerned on the response of real export to involved variables in the model. Generally, figure 2 and table 4 show similar result as what is shown by the long-run term impact. In detail, real exports respond positively to GDP and bilateral real exchange rate of IDR against US\$. Since the Marshall-Lerner condition failed to be existed, the J-Curve phenomenon cannot be framed. In contrast, it reacts negatively to bilateral real exchange rate volatility of IDR against US\$ and bilateral real exchange rate of RMB against US\$. These results conforms the variance decomposition where the effects are mostly positive and small (Table 4).

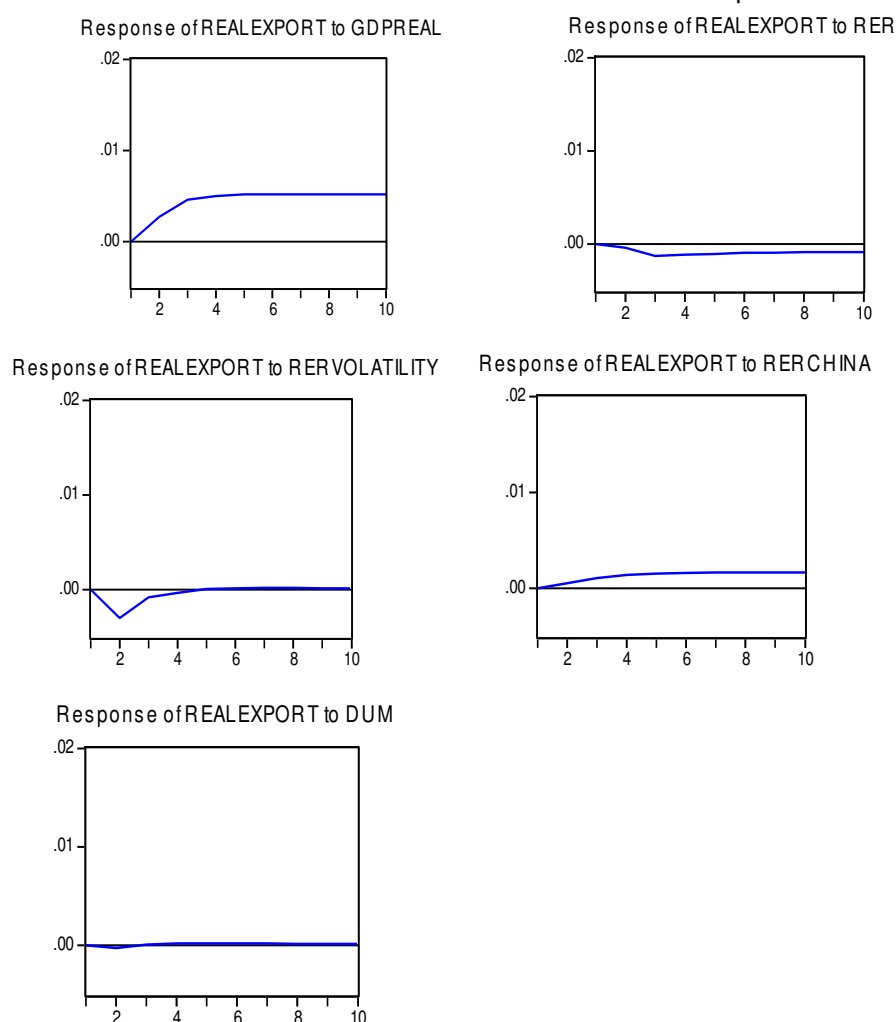


Figure 2. Impulse Response Function
Source: Authors' Calculations

Table 4:Variance Decomposition

Period	S.E.	REALEXPORT	GDPREAL	RER	RERVOL	RERC	DUM
1	0.016	100.00	0.000	0.000	0.000	0.000	0.000
3	0.029	95.219	3.297	0.204	1.107	0.165	0.008
5	0.040	93.527	5.203	0.261	0.623	0.374	0.011
7	0.048	92.803	5.998	0.253	0.432	0.502	0.012
9	0.055	92.436	6.400	0.244	0.331	0.578	0.011

Note: Cholesky Ordering: REALEXPORT, GDPREAL, RER,RERVOL, RERC, DUM

As pointed out by Enders (1995), the forecast error variance decomposition tells the proportion of the movements in a sequence because its own shock against other variables shock. In the other words, the variance decomposition is a calculation of variance forecast error in the percentage framed by another variable within the short run dynamics and interactions. Table 4 presents the result of variance decomposition within 1, 3, 5, 7, and 9 periods. It confirms that the real exports are mostly related to its own variation from 95.2 percent in period 3 up to 92.4 percent in period 9. The highest contributor variable on real exports changes is GDP, meanwhile, variable RER, RERVOL, RERC, and DUM shows small effects on the dynamic of real exports.

Research Analysis and Policy Recommendations

Based on the findings, first, we can suggest that Indonesia should encourage their domestic market rather than international market. Krugman and Obstfeld (2006) summarized that if there is no domination of the volume effect over the price effect, hence the phenomenon of Marshall-Lerner condition will never be satisfied. Thus, since this study unable to satisfy both of the Marshall-Lerner condition and J-Curve phenomenon, it implies that the depreciation of IDR RER might not improve the trade balance of Indonesia. In other words, the IDR depreciation might only affect positively from volume side, but without improvement in the exports value. It also can be inferred that the products of Indonesia have unprepared well yet to be exports product. Therefore, this study suggests to change the orientation from international to domestic market.

Second, comparing the values of estimated coefficient of RER and RERC, it implies that the depreciation of IDR will increase the exports of Indonesia to the US (0.32 percent). However, it yields smaller amount than if RMB is depreciated against the US\$ (0.36 percent). In this case, the relationship between these two products is substitution, hence the depreciation of RMB will suffer Indonesian exports to the US. In the other words, the product of Indonesia is less competitive than the product from China in the US market. Therefore, this study propose to strengthen the establishment of real sectors in order to improve the efficiency in production cost before to promote “go international”.

Third, since the Chinese government has made a conscious decision to embark on far-reaching economic liberalization and enlisted as a new member of World Trade Organization (WTO) in 2001, the exports, Foreign Direct Investment (FDI) and overall growth prospects of China have significantly increased (Rawski, 2009; Yamazawa and Imai, 2001). Data from the UN COMTRADE shows that Chinese exports have achieved a record i.e. exceeding Japan's exports since 2004. In 1988 the number of FTCs reached more than 5.000 with full authority in trade (Panagariya, 2006). Considering these Chinese hegemony in the Asia, in the short run the government of Indonesia should determine the variety of exports goods so that the exports product become complementary with China's product. However, in the long run, the government should develop genuine products that can be competitive advantage for Indonesia.

CONCLUSIONS

This study finds that the RMB depreciation has negative impacts on the exports of the Indonesia to the US. It indicates that the relationship between Chinese and Indonesian exports to the US market is substitution. It also confirms that there no Marshall-Lerner condition indicating the J-curve phenomenon does not exist.

There are three policy recommendations regarding the empirical results, as follows, (a) in short run, it is recommended to change the orientation from international to domestic market, (b) before the government of Indonesia promote “go international”, it is suggested to induce and support the real sectors so that there will be an improvement in

the efficiency cost of production, and (c) in the long run, Indonesian government should develop competitive products in international market.

Finally, this study focuses on the impact of Chinese Renminbi on the bilateral exports between Indonesia to US in the long run. The short run, equation from import side and sectoral (commodity) level analyses are excluded. For future research, it is encouraged to explore the impact of Chinese renminbi by considering from two sides (export and import) and commodity level in order to get deeper and stronger empirical analysis both in the short run and the long run.

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Appendix

The Formulas

Real Exports

$$EX_{ijt} = \ln \left(\frac{NIM_{ijt}}{IMUV_{jt}} \right) \times 100 \quad (4)$$

Due to that the data of export unit value from Indonesia are incomplete, the real exports are calculated as of in the study of Baak (2006). EX_{ijt} equals to the logarithm value of the quarterly nominal imports, NIM_{ijt} , of country j from i divided by import unit value index of a country i ($IMUV_{it}$) times 100.

Real Exchange Rate Volatility

The $RERVOL_{ijt}$ represents the logarithm value of real exchange rate volatility. It is measured as natural logarithm of the absolute quarterly standard deviation of monthly real exchange rate which formulated as follows²:

$$RERVOL_{ijt} = \ln \left[\sqrt{\frac{1}{n-1} \sum_{k=tm}^{tn} [RER_{ijk} - \overline{RER}_{ij}]^2} \right] \quad (3)$$

RER_{ijk} represents the monthly real exchange rate, meanwhile \overline{RER}_{ij} is the quarterly average of monthly real exchange rates and k is the index of the months in a quarter.

² The equation of the logarithm value of real exchange rate volatility above is a correction of formula which has been employed by Baak (2007), as follows:

$$RERVOL_{ijt} = \ln \left(\sqrt{\frac{1}{n-1} \sum_{k=tm}^{tn} [RER_{ijk} - \overline{RER}_{ij}]^2} \right)$$

According to the equation of Baak (2007), there is a possibility to compute negative and/or zero value from standard deviation, which is impossible to be calculated. Thus, we argue that absolute brackets needs to be added to correct in Baak's formula.

Lag Length Criteria

Lag	LogL	LR	FPE	AIC	SC	HQ
0	136.2622	NA	2.09e-09	-2.960504	-2.791594	-2.892454
1	798.3695	1218.879	1.38e-15*	-17.19022*	-16.00785*	-16.71387*
2	832.9609	58.96270*	1.44e-15	-17.15820	-14.96238	-16.27356
3	852.1160	30.03860	2.18e-15	-16.77536	-13.56609	-15.48243
4	875.9133	34.07350	3.03e-15	-16.49803	-12.27530	-14.79680
5	896.5000	26.66913	4.73e-15	-16.14773	-10.91154	-14.03820
6	913.3985	19.58690	8.43e-15	-15.71360	-9.463958	-13.19578
7	947.4973	34.87373	1.09e-14	-15.67039	-8.407292	-12.74427
8	973.1739	22.75882	1.88e-14	-15.43577	-7.159214	-12.10135
9	1008.806	26.72403	2.94e-14	-15.42741	-6.137395	-11.68470

*indicates lag order selected by the criterion

LR: sequential modified LR test statistic (each test at 5% level)

FPE: Final prediction error

AIC: Akaike information criterion

SC: Schwarz information criterion

HQ: Hannan-Quinn information criterion